6. miniworkshop difrakce a ultraperiferálních srážek

Fitting with correlated uncertainties

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R. J. Barlow, Practical statistics for particle physics <u>https://e-publishing.cern.ch/index.php/CYRSP/article/view/1384</u>

Suppose that you know the probability density function P(x; a) for some process and you have a data set $\{x_i\}$

What is the best value of *a*?

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There may be many parameters

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Parameter estimation

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Parameter estimation

The data may be multidimensional

Suppose that you know the probability density function P(x; a) for some process and you have a data set $\{x_i\}$

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An estimator $\hat{a}(x_1, ..., x_n)$ is a function of the data that gives a value for the parameter a

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An ideal estimator would

- Consistent. $\hat{a}(x_1$
- Unbiased: $\langle \hat{a} \rangle =$
- Efficient: $\langle (\hat{a} \hat{a}) \rangle$
- Invariant: $\hat{f}(a)$ =

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The data may be multidimensional

be

$$(\dots x_N) \to a \text{ as } N \to \infty,$$

 $(= a, a)^2$ is as small as possible,
 $= f(\hat{a}).$

Suppose that you know the probability density function P(x; a) for some process and you have a data set $\{x_i\}$

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There is no such a thing	 An ideal estimator would k
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An estimator $\hat{a}(x_1, \dots, x_n)$ is a function of the data that gives a value for the parameter *a*

be

$$(\dots x_N) \to a \text{ as } N \to \infty,$$

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 $= f(\hat{a}).$

The maximum likelihood (ML) estimator adjusts *a* so that the likelihood of the sample is maximised

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ML is

- It is consistent.
- It is biased, but bias falls like 1/N.
- It is efficient for the large N.

R. J. Barlow, Practical statistics for particle physics https://e-publishing.cern.ch/index.php/CYRSP/article/view/1384

The maximum likelihood (ML) estimator adjusts a so that the likelihood of the sample is maximised

- It is invariant—doesn't matter if you reparametrize a.

If you have Gaussian measurements of y taken at various x values with a measurement error σ , and a prediction y = f(x; a) then

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Gaussian probabilities

$$P(y; x, a) = \frac{1}{\sigma\sqrt{2\pi}} \exp^{-\frac{(y - f(x; a))^2}{2\sigma^2}}$$

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To maximise $\ln L$ is equivalent to minimise:

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If you have Gaussian measurements of y taken at various x values with a measurement error σ , and a prediction y = f(x; a) then

$$\blacktriangleright \ln L = -\sum \frac{(y - f(x; a))^2}{2\sigma^2} + \text{constants}$$

$$\chi^2$$
 minimisation
 $\chi^2 = \sum \frac{(y - f(x; a))^2}{\sigma^2}$

Gaussian probabilities

$$P(y; x, a) = \frac{1}{\sigma\sqrt{2\pi}} \exp^{-\frac{(y - f(x; a))^2}{2\sigma^2}}$$

To maximise $\ln L$ is equivalent to minimise:

In case of a histogram with *n* entries, the uncertainty, given by Poisson statistics, is $\sigma^2 = n$ which breaks down for empty bins

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$$\blacktriangleright \quad \ln L = -\sum \frac{(y - f(x; a))^2}{2\sigma^2} + \text{constants}$$

$$\chi^{2} \text{ minimisation}$$

$$\chi^{2} = \sum \frac{(y - f(x; a))^{2}}{\sigma^{2}}$$

In mathematical optimization, the method of Lagrange multipliers is a strategy for finding the local maxima and minima of a function subject to equation constraints (i.e., subject to the condition that one or more equations have to be satisfied exactly by the chosen values of the variables).^[1] It is named after the mathematician Joseph-Louis Lagrange. The basic idea is to convert a constrained problem into a form such that the derivative test of an unconstrained problem can still be applied. The relationship between the gradient of the function and gradients of the constraints rather naturally leads to a reformulation of the original problem, known as the Lagrangian function.^[2]

form the Lagrangian function,

$$\mathcal{L}(x,\lambda)\equiv f(x)+\lambda\cdot g(x)$$

and find the stationary points of $\mathcal L$ considered as a function of x and the Lagrange multiplier λ . This means that all partial derivatives should be zero, including the partial derivative with respect to λ .^[3]

$$rac{\partial \mathcal{L}}{\partial x} = 0 \qquad ext{and} \qquad rac{\partial \mathcal{L}}{\partial \lambda} = 0 \ ;$$

or equivalently

$$rac{\partial f(x)}{\partial x} + \lambda \cdot rac{\partial g(x)}{\partial x} = 0 \qquad ext{and} \qquad g(x) = 0 \;.$$

the stationary points from the definiteness of the bordered Hessian matrix.^[6]

https://en.wikipedia.org/wiki/Lagrange multiplier

Lagrange multipliers

The method can be summarized as follows: In order to find the maximum or minimum of a function f(x) subjected to the equality constraint g(x) = 0,

The solution corresponding to the original constrained optimization is always a saddle point of the Lagrangian function, [4][5] which can be identified among

$$\chi_{\exp}^2(\boldsymbol{m}, \boldsymbol{b}) = \sum_i \frac{[m^i - \sum_j]^i}{\Delta}$$

Measurement of the Inclusive ep Scattering Cross Section at Low Q^2 and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 <u>https://inspirehep.net/record/817368?ln=en</u>

 $\frac{\Gamma_{j}^{i}b_{j} - \mu^{i}]^{2}}{\Delta_{i}^{2}} + \sum_{j} b_{j}^{2}.$ (27)

Theory

$$\chi^{2}_{\exp}(\boldsymbol{m}, \boldsymbol{b}) = \sum_{i} \frac{[m^{i} - \sum_{j}]^{i}}{\Delta}$$

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 $\frac{\Gamma_j^i b_j - \mu^i]^2}{\Lambda_{\cdot}^2} + \sum_j b_j^2. \qquad (27)$

Theory

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$$\chi^{2}_{\exp}(\boldsymbol{m}, \boldsymbol{b}) = \sum_{i} \frac{[m^{i} - \sum_{j}]_{i}}{2}$$

Statistical and uncorrelated uncertainty

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This χ^2 prescription is suitable for measurements in which the uncertainties are absolute, i.e. do not depend on the central value of the measurement.



Theory

$$\chi^{2}_{\exp}(\boldsymbol{m}, \boldsymbol{b}) = \sum_{i} \frac{[m^{i} - \sum_{j}]_{i}}{2}$$

Statistical and uncorrelated uncertainty

But correlated and uncorrelated systematic errors are frequently, to a good approximation, proportional to the central values (multiplicative errors)

Measurement of the Inclusive ep Scattering Cross Section at Low Q² and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 https://inspirehep.net/record/817368?ln=en



This χ^2 prescription is suitable for measurements in which the uncertainties are absolute, i.e. do not depend on the central value of the measurement.



$$\chi_{\exp}^{2}(\boldsymbol{m}, \boldsymbol{b}) = \sum_{i} \frac{[m^{i} - \sum_{j} \gamma_{j}^{i} m^{i} b_{j} - \mu^{i}]^{2}}{\delta_{i, \text{stat}}^{2} \mu^{i} (m^{i} - \sum_{j} \gamma_{j}^{i} m^{i} b_{j}) + (\delta_{i, \text{uncor}} m^{i})^{2}} + \sum_{j} b_{j}^{2}.$$
(31)

Measurement of the Inclusive ep Scattering Cross Section at Low Q^2 and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 https://inspirehep.net/record/817368?ln=en

$$\chi_{\exp}^{2}(\boldsymbol{m}, \boldsymbol{b}) = \sum_{i} \frac{[m^{i} - \sum_{j} \gamma_{j}^{i} m^{i} b_{j} - \mu^{i}]^{2}}{\delta_{i, \text{stat}}^{2} \mu^{i} (m^{i} - \sum_{j} \gamma_{j}^{i} m^{i} b_{j}) + (\delta_{i, \text{uncor}} m^{i})^{2}} + \sum_{j} b_{j}^{2}.$$

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Measurement of the Inclusive ep Scattering Cross Section at Low Q^2 and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 <u>https://inspirehep.net/record/817368?ln=en</u> Relative correlated uncertainty source *j* in measurement *i*

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Measurement of the Inclusive ep Scattering Cross Section at Low Q^2 and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 <u>https://inspirehep.net/record/817368?ln=en</u> Relative correlated uncertainty source *j* in measurement *i*

Relative statistical uncertainty

$$\begin{aligned}
\chi^{2}_{exp}(\boldsymbol{m}, \boldsymbol{b}) \\
= \sum_{i} \frac{[m^{i} - \sum_{j}]}{\delta_{i, stat}^{2} \mu^{i} (m^{i} - \sum_{j}]} \\
+ \sum_{j} b_{j}^{2}.
\end{aligned}$$

Measurement of the Inclusive ep Scattering Cross Section at Low Q^2 and x at HERA H1 Collaboration, Eur.Phys.J. C63 (2009) 625-678 https://inspirehep.net/record/817368?ln=en





This is the formula we will use



[1] Energy dependence of exclusive \$J/\psi\$ photoproduction off protons in ultraperipheral p-Pb collisions at $\sqrt{s_{NN}} = 5.02 \text{ TeV}$ ALICE Collaboration, Eur.Phys.J. C79 (2019) no.5, 402 https://inspirehep.net/record/1693305?ln=en

Example: power law



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Example: power law

Generate cross sections using realistic uncorrelated uncertainties





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Example: power law

Generate cross sections using realistic uncorrelated uncertainties



Overall normalisation (eg. luminosity or BR) Affecting only 'forward' data Affecting only 'semi-forward' data Affecting only 'central-barrel' data

```
// parameters for ideal power law
// _____
double lambdaPL = 0.7;
double nPL = 70; // nb
double wPL = 90; // GeV
// info to generate power law
const int nPoints = 9; // number of measurements to be simulated
const double wPoints[nPoints] = {24, 30, 40, 50, 70, 130, 190, 390, 700}; // GeV
const double staUnc[nPoints] = {0.05, 0.03, 0.05, 0.07, 0.07, 0.03, 0.03, 0.07,0.04}; // rel sta
const double uncUnc[nPoints] = {0.08, 0.08, 0.08, 0.07,
                         0.05, 0.07, 0.07, 0.07, 0.10}; // uncorrelated unceratainty in
const int nCor = 4; // number of correlated sources of uncertainty
const double normUnc = 0.03; // global normalization correlated error
const double fwdUnc = 0.05; // correlated uncertainty for fwd points
const double sfUnc = 0.04; // correlated uncertainty for semi-fwd points
const double cbUnc = 0.04; // correlated uncertainty for central barrel points
const double corUnc[nCor][nPoints] = // relative correlated uncertainty
 {
   normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc,
   fwdUnc, fwdUnc, fwdUnc, 0.0, 0.0, 0.0, 0.0, 0.0, fwdUnc, // only fwd
   0.0, 0.0, 0.0, sfUnc, sfUnc, 0.0, 0.0, sfUnc, 0.0, // only semifwd
   0.0, 0.0, 0.0, 0.0, 0.0, cbUnc, cbUnc, 0.0, 0.0 // only semifwd
```

Snippets (1/2)

at unc	
percent	

```
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                              0.05, 0.07, 0.07, 0.07, 0.10}; // uncorrelated unceratainty in percent
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const double corUnc[nCor][nPoints] = // relative correlated uncertainty
   normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc, normUnc,
   fwdUnc, fwdUnc, fwdUnc, 0.0, 0.0, 0.0, 0.0, 0.0, fwdUnc, // only fwd
    0.0, 0.0, 0.0, sfUnc, sfUnc, 0.0, 0.0, sfUnc, 0.0, // only semifwd
    0.0, 0.0, 0.0, 0.0, 0.0, cbUnc, cbUnc, 0.0, 0.0 // only semifwd
```

Snippets (1/2)

```
aenerated power law
double csGen[nPoints]; // nb
double csPL[nPoints]; // nb
double suUnc[nPoints]; // stat+unc uncertaity in nb
void generateCS()
 // initialise random number generator
 gRandom->SetSeed(0);
 cout << endl << " Random seed: " << gRandom->GetSeed() << endl;</pre>
 // generate cross section
 for(int i=0; i<nPoints;i++) {</pre>
   csPL[i] = nPL*TMath::Power(wPoints[i]/wPL,lambdaPL);
   csGen[i] = csPL[i]*gRandom->Gaus(1,staUnc[i]);
   csGen[i] *= gRandom->Gaus(1,uncUnc[i]);
   suUnc[i] = csGen[i]*TMath::Sqrt(staUnc[i]*staUnc[i]+uncUnc[i]*uncUnc[i]);
 // shift full cross section
  double normCS = gRandom->Gaus(1,normUnc);
 for(int i=0; i<nPoints;i++) csGen[i] *= normCS;</pre>
  cout << " normalisation shift: " << normCS << endl;</pre>
 // shift fwd cross sections
 double normFW = gRandom->Gaus(1,fwdUnc);
 csGen[0] *= normFW; csGen[1] *= normFW; csGen[2] *= normFW; csGen[8] *= normFW;
 cout << " fwd shift: " << normFW << endl;</pre>
 // shift sf cross sections
 double normSF = gRandom->Gaus(1,sfUnc);
 csGen[3] *= normSF; csGen[4] *= normSF; csGen[7] *= normSF;
 cout << " semi-fwd shift: " << normSF << endl;</pre>
 // shift cb cross sections
  double normCB = gRandom->Gaus(1,cbUnc);
 csGen[5] *= normCB; csGen[6] *= normCB;
 cout << " central barrel shift: " << normCB << endl;</pre>
 cout << endl;</pre>
```



```
// chi2 function to fit data according to prescription in
// Eur.Phys.J. C63 (2009) 625-678, section 9, eq (31)
void fcnChi2Model(int &npar, double *gin, double &f, double *par, int iflag)
 // use unused parguments to avoid compiler message
 double tmp = npar; tmp = gin[0]; tmp = iflag;
 // define model according to:
 // ch2 = sum_i [m_i-mu_i-Sij]^2/D + Sbj
 // Sij = sum_j g_ij*m_i*b_j
 // Sbj = sum_j (b_j)^2
 // D = d_i_stat^2*mu_i*(m_i-Sij)+(d_i,unc*m_i)^2
 // mu_i is the measurement
 // g_ij relative normalization uncertainty at point i from source j
 // d_i = relative uncertainty (either stat or uncorr)
 double bj[nCor] = {par[2],par[3],par[4],par[5]};
 double Sbj = 0;
 for(int j=0;j<nCor;j++) Sbj += (bj[j]*bj[j]);</pre>
 double chi2 = 0;
 for (int i=0; i<nPoints;i++) {</pre>
   double mu_i = csGen[i];
   double m_i = par[0]*TMath::Power(wPoints[i]/wPL,par[1]);
   double Sij = 0;
   for(int j=0;j<nCor;j++) Sij += m_i*bj[j]*corUnc[j][i];</pre>
   double d_stat_i = staUnc[i];
   double d_unc_i = uncUnc[i];
   double D = (d_stat_i*d_stat_i*mu_i*(m_i-Sij))+((d_unc_i*m_i)*(d_unc_i*m_i));
    chi2 += ((m_i-mu_i-Sij)*(m_i-mu_i-Sij)/D);
 chi2 += Sbj;
 f = chi2;
 gChi2 = f;
```



```
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   double m_i = par[0]*TMath::Power(wPoints[i]/wPL,par[1]);
   double Sij = 0;
   for(int j=0;j<nCor;j++) Sij += m_i*bj[j]*corUnc[j][i];</pre>
   double d_stat_i = staUnc[i];
   double d_unc_i = uncUnc[i];
   double D = (d_stat_i*d_stat_i*mu_i*(m_i-Sij))+((d_unc_i*m_i)*(d_unc_i*m_i));
   chi2 += ((m_i-mu_i-Sij)*(m_i-mu_i-Sij)/D);
 chi2 += Sbj;
 f = chi2;
 gChi2 = f;
```

Snippets (2/2)

```
void powerLawFit()
 // generate data set
 generate(S();
 // now do TMinut
 cout << endl<< endl<< endl;</pre>
 cout << endl<< endl<< endl;</pre>
 // initialize minuit with a maximum of parameters
 const int nPar = 2+nCor;
 TMinuit *myMinuit = new TMinuit(nPar);
 // set the function with the minimization process
 myMinuit->SetFCN(fcnChi2Model);
 // define parameters
 double minBj = -10;
 double maxBj = -minBj;
 myMinuit->DefineParameter(0, "N", 70, 1, 0.0, 1000.);
 myMinuit->DefineParameter(1, "#delta", 0.7, 0.1, 0.1, 10.);
 myMinuit->DefineParameter(2,"global norm",0.001,0.0001,minBj,maxBj);
 myMinuit->DefineParameter(3,"fwd norm",0.001,0.0001,minBj,maxBj);
 myMinuit->DefineParameter(4,"sf norm",0.001,0.0001,minBj,maxBj);
 myMinuit->DefineParameter(5,"cb norm",0.001,0.0001,minBj,maxBj);
 // migrad
```

myMinuit->SetMaxIterations(500); myMinuit->Migrad();

______ // get results



Random seed: 2584683516 normalisation shift: 0.954077 fwd shift: 0.985996 semi-fwd shift: 0.957315 central barrel shift: 1.00353

Test 1

Random seed: 2584683516 normalisation shift: 0.954077 fwd shift: 0.985996 semi-fwd shift: 0.957315 central barrel shift: 1.00353



Random seed: 2584683516 normalisation shift: 0.954077 fwd shift: 0.985996 semi-fwd shift: 0.957315 central barrel shift: 1.00353



Test 1

Correlation matrix



Random seed: 3900432653 normalisation shift: 0.986186 fwd shift: 1.00663 semi-fwd shift: 1.04337 central barrel shift: 1.01852





12

0.8

-0.4

- 0.2

—––0.;

Random seed: 3900432653 normalisation shift: 0.986186 fwd shift: 1.00663 semi-fwd shift: 1.04337 central barrel shift: 1.01852



Test 2

Correlation matrix



).8

).4

).2)

-0.:

Code tested, and attached to the indico page

Formalism to estimate parameters when there are correlated uncertainties